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## RESEARCH ARTICLE

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### Key Points:

- Compared with concentration-independent factory-specified errors, field characterization of concentration-dependent errors of isotope ratio infrared spectroscopy instruments improves parameter estimation of the York's solution regression
- Good agreement is achieved among the flux-gradient, the Keeling plot method using York's solution, and the Keeling plot method using ordinary least squares
- The Keeling plot method gives better results by using measurements made at two heights than at only one height

### Supporting Information:

- Supporting Information S1

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## Determining the Isotopic Composition of Surface Water Vapor Flux From High-Frequency Observations Using Flux-Gradient and Keeling Plot Methods

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**Abstract** The isotopic composition of surface water vapor flux ( $\delta_E$ ) is a quantity frequently used to investigate the local and regional water cycle. This study reports the results of a comparative evaluation of  $\delta_E$  determined with the Keeling plot and the flux-gradient methods using high-frequency data collected at a cropland site and a lake site. Three regression models, ordinary least squares (OLS), York's solution (YS), and geometric mean regression, were tested with the Keeling plot method. Results show that concentration-dependent field characterization of measurement errors can improve the estimation of the YS regression. For both sites, broad agreement was achieved among the Keeling plot method with YS regression, the Keeling plot method with OLS regression and the flux-gradient method. For the lake site, OLS was the least biased of the three regression models in reference to the  $\delta_E$  calculated by the Craig-Gordon (CG) model of isotopic evaporation of open water. Good agreement was also achieved between the flux-gradient method and the CG model at the lake site under open-fetch conditions. A footprint analysis suggests that the Keeling method with OLS regression may be less sensitive to fetch than the flux-gradient method.

## 1. Introduction

The isotopic composition of surface water vapor flux ( $\delta_E$ ) is a key parameter in studies of the water cycle using isotopic tracer methods. It is used for estimating lake evaporation (Gibson et al., 1993; Xiao et al., 2017), constraining local moisture recycling (Bowen et al., 2019; Gat et al., 1994; Griffis et al., 2016; S. Wang et al., 2016; Xiao et al., 2018), characterizing sources of moisture in the atmospheric boundary layer (Lee et al., 2007; Simonin et al., 2014; Welp et al., 2008, 2012; Zannoni, Steen-Larsen, Stenni, et al., 2019), and partitioning of evapotranspiration in ecosystems (Good et al., 2014, 2015; Lu et al., 2017; Sun et al., 2019; Wei et al., 2018; Wen et al., 2016). The Keeling plot method and the flux-gradient method are two common measurement strategies for determining  $\delta_E$ . Each strategy requires certain conditions about atmospheric mixing near the ground. The Keeling plot method was originally developed for CO<sub>2</sub>. By extending it to water vapor, it assumes (1) that surface evaporation is solely responsible for observed variations in the isotopic composition of water vapor  $\delta_v$  and in the water vapor concentration  $c$  and (2) that  $\delta_E$  remains invariant during the observational period. If the two underlying assumptions are satisfied, the mixing of the evaporated vapor with vapor in the surface-layer air can be described by (X. Wang & Yakir, 2000; Yopez et al., 2003)

$$\delta_v = a + b(1/c) \quad (1)$$

where  $a = \delta_E$  and  $b = c_0(\delta_0 - \delta_E)$  are intercept and slope coefficients, respectively, with  $c_0$  and  $\delta_0$  denoting the background concentration and vapor delta value.

The flux-gradient method determines  $\delta_E$  from the ratio of the vertical concentration gradient of the minor to that of the major isotopologue (e.g., Lee et al., 2007). The molar ratio of the H<sub>2</sub><sup>18</sup>O flux to the H<sub>2</sub><sup>16</sup>O flux is given by

$$R_E = (c_2^i - c_1^i) / (c_2 - c_1) \quad (2)$$

where  $c^i$  and  $c$  denote the mean molar mixing ratio of  $\text{H}_2^{18}\text{O}$  and  $\text{H}_2^{16}\text{O}$  of an observation period, respectively, and subscripts 1 and 2 denote the upper and the lower measurement level, respectively. The molar flux ratio  $R_E$  is then converted to the  $\delta$  scale to give  $\delta_E$ . The flux-gradient method assumes that the diffusion of the  $\text{H}_2^{18}\text{O}$  and  $\text{H}_2^{16}\text{O}$  molecules is identically efficient in the atmospheric surface layer so that the diffusivity coefficient cancels out when performing the flux ratio calculation (Griffis et al., 2005). Equation 2 is used when measurements are made at two heights. In some applications where measurements are made at more than two heights,  $R_E$  can be obtained by linear regression of  $c^i$  against  $c$ . Another implicit assumption is that the footprint of measurement at the upper level and that at the lower level lie in the same source area (Griffis et al., 2007). In situations where the two measurement heights are far apart vertically or where the fetch of the target surface is limited, the evaporation of a source upwind of the target surface can “contaminate” the upper measurement more than the lower measurement, causing errors in the measured flux and the flux ratio (Horst, 1999).

Before isotope ratio infrared spectroscopy (IRIS) instruments became available, application of the Keeling plot method involved measurement of water vapor concentration,  $c$ , and collection of water vapor via cold traps for determination of  $\delta_v$ . In order to obtain enough trapped samples for the regression analysis, the observation period often extended several hours or longer (Delattre et al., 2015; Yopez et al., 2003, 2005; Zannoni, Steen-Larsen, Rampazzo, et al., 2019). However, temporal changes in atmospheric forcing, such as relative humidity, and cloudiness, can cause large short-term (minutes to hours) fluctuations in  $\delta_E$  of land evapotranspiration (Dubbert & Werner, 2019; Good et al., 2012; Lee et al., 2007; Quade et al., 2019; Welp et al., 2008; Wen et al., 2016) and open-water evaporation (Xiao et al., 2017). One consequence of  $\delta_E$  variations is that  $\delta_v$  may no longer be linear with  $1/c$  over longer integration intervals. When this occurs, the validity of the second Keeling plot assumption, that  $\delta_E$  remains invariant during the observational period, is questionable (Pataki et al., 2003). Furthermore, if the observation period is too long, temporal changes in  $\delta_v$  and  $c$  can result from mesoscale and synoptic-scale atmospheric events unrelated to surface evaporation, even at a measurement height very close to the surface, raising doubt about the first Keeling plot assumption, that surface evaporation is solely responsible for observed variations (Lee et al., 2006). Obviously, when one or both of the Keeling plot method assumptions are violated, the intercept coefficient  $a$  is no longer a true representation of  $\delta_E$ , regardless of which regression model is used for parameter estimation.

With the IRIS technology, it is possible to apply the Keeling plot method to the high-frequency  $c$  and  $\delta_v$  time series data collected in short observation periods (e.g., 1 h). The idea of applying the Keeling plot method to high frequency time series was first proposed by Bowling et al. (1999) before the emergence of the IRIS technology and was later tested with  $\text{CO}_2$  isotope data collected with an IRIS instrument (Griffis et al., 2004). Because the data are collected at a high frequency, the sample size is large (typically > 700 in 1 h), effectively increasing the variability in the observed water vapor concentration and decreasing the uncertainty of parameter estimation (Good et al., 2012). By restricting the regression based on the high-frequency data to a short period (hourly), errors arising from the two underlying assumptions should be small. Therefore, it is possible to determine  $\delta_E$  using the same data with either the Keeling plot method or the flux-gradient method. A practical question is whether these two methods agree with each other under field conditions.

The Keeling plot results are sensitive to how measurement errors are specified. The ordinary least squares (OLS) regression model, the most common model for estimating the regression coefficients, assumes that all measurement errors occur in the dependent variable ( $\delta_v$ , Equation 1) and no errors exist in the independent variable (concentration,  $c$ ). However, the concentration data can also suffer from measurement errors, although they are generally smaller than errors in isotope data. To account for errors in both, some researchers recommend that the geometric mean regression (GMR) model be used for parameter estimation (e.g., Ogée et al., 2003; Pataki et al., 2003). The GMR model assumes that the normalized error in  $x$  is equal to the normalized error in  $y$ . More recently, Wehr and Saleska (2017) recommend that a general regression model, named here the York’s solution (YS) model, be used with the Keeling plot method. The YS

model takes into account error structures of  $x$  and  $y$  separately and also the correlation between these errors. In Wehr and Saleska's evaluation of the YS model, errors in the dependent and independent variables are concentration-independent instrument precisions. In the case of water vapor isotope, the  $\delta_v$  precision is generally dependent on the water vapor concentration (Salmon et al., 2019; Sturm & Knohl, 2010). Because this concentration dependence is often instrument-specific (even for the same instrument brand), error characterization in field conditions may help improve YS parameter estimation. Another reason for field characterization of instrument errors is that using high-frequency data may lead to strongly correlated errors in  $x$  and in  $y$ , as discussed by Wehr and Saleska (2017).

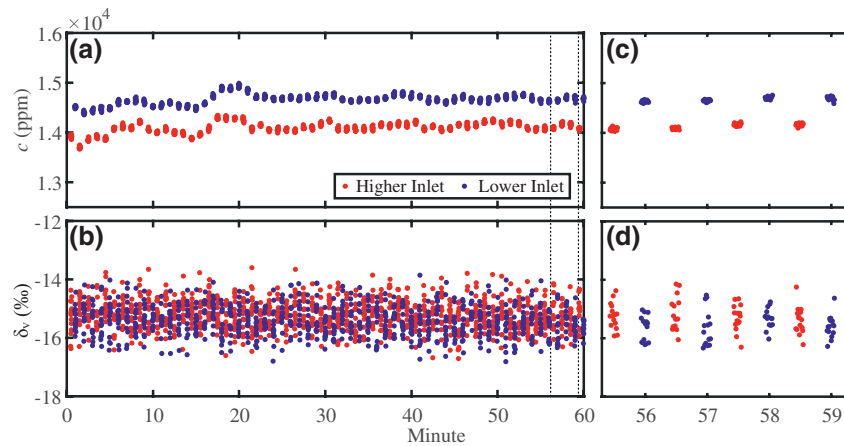
Another practical issue concerns bias errors of the Keeling plot and the flux-gradient methods due to non-ideal experimental conditions. As mentioned earlier, in the case of the Keeling plot method, bias errors can occur even with a perfect statistical model, if the underlying assumptions about atmospheric mixing are not met. Assessment of systematic biases is challenging for land ecosystems because the true  $\delta_E$  is not known *a priori*, even in isotopic steady state. In isotopic steady state, the isotopic composition of plant transpiration approaches that of the xylem water which is a measurable quantity, but  $\delta_E$  of evapotranspiration is also influenced by soil evaporation whose isotopic composition is generally unknown (Yakir & daSternberg, 2000). To overcome a similar problem for  $\text{CO}_2$ , some researchers used synthetic data that combines a hypothetical flux isotopic signal with Monte-Carlo type random variations superimposed on the concentration and the  $\delta$  variable (Chen et al., 2017; Kayler et al., 2010; Vardag et al., 2016; Wehr & Saleska, 2017; Zobitz et al., 2006). This synthetic approach is less feasible for water vapor because it is difficult to assign a representative value for  $\delta_E$  due to its high temporal variability. According to in situ observations of Welp et al. (2008),  $\delta_E$  can vary by as much as 40‰ in the course of a day. To overcome this drawback, we tested a new strategy to assess bias errors, using data collected at a lake site. Here, the benchmark is the  $\delta_E$  calculated with the Craig-Gordon (CG) model of the isotopic composition of open-water evaporation (Craig & Gordon, 1965). Because the model is grounded on well-established principles of equilibrium and kinetic fractionation of open-water evaporation (Gonfiantini et al., 2018), it can provide an independent estimate of  $\delta_E$  for evaluating the Keeling plot and the flux-gradient methods.

In this study, we report the results of a comparative evaluation of the Keeling plot method and the flux-gradient method using high-frequency data collected with IRIS instruments at a cropland site and a lake site. Wehr and Saleska (2017) have conducted a comprehensive evaluation of the Keeling plot method for  $\text{CO}_2$ . The analysis presented below can be viewed as a test of their findings for water vapor. The specific objectives of this study are (1) to determine if concentration-dependent field characterization of error structures of the IRIS instruments can improve parameter estimation of the YS regression, (2) to characterize the relative agreement between the flux-gradient method and the Keeling plot method with three regression models (OLS, GMR, and YS), and (3) to evaluate bias errors of the Keeling plot method and the flux-gradient method against the CG model prediction. Even though hydrogen isotopes were also measured in the two experiments, we restrict our analysis to oxygen isotopes.

## 2. Materials and Methods

### 2.1. Sites and Instruments

The data sets used in this study were obtained in two field experiments. The first experiment was conducted in an irrigated maize field in Zhangye, Gansu Province, in Northwest China (38° 51' N, 100° 22' E) in 2012 (Wen et al., 2016). The fetch was greater than 200 m in all the directions. The  $\text{H}_2\text{O}$ , HDO, and  $\text{H}_2^{18}\text{O}$  concentrations were measured at two heights (0.5 and 1.5 m) above the canopy top, with an IRIS water vapor isotope analyzer (Model L1102-i, Picarro Inc., CA, USA) at 0.2 Hz. The canopy height varied from 0.1 m at the beginning to 1.6 m at end of the experiment. The analyzer was customized to improve its time response. Switching between the two intake heights occurred every 2 min and the measurement became stable after 25 s (5 datapoints; Wen et al., 2016, their Figure 1). The last eight datapoints after each switching, corresponding to the last 40 s, were used in the analysis. The analyzer was calibrated in situ with a liquid vaporization module (Picarro Inc.) and a CTC Analytics Prep and Load liquid autosampler (LEAP Technologies, Carrboro, NC, USA) using a single liquid water standard with a  $\delta^{18}\text{O}$  value of  $-14.29\%$ . There were three concentrations of calibration vapor, each measured for 25 min. After each calibration, 3 h



**Figure 1.** Temporal variation of H<sub>2</sub>O mixing ratio (a) and  $\delta_v$  (b) at the lower inlet (blue dots) and higher inlet (red dots) at Lake Taihu between 16:00 and 17:00 local time on October 22, 2014. Panels (c) and (d) are the corresponding zoom-in plot of the dotted box in panels (a) and (b).

were spent on the measurement of ambient air. A linear interpolation between two consecutive calibration cycles was used to obtain the span for correcting the ambient air measurements (Huang & Wen, 2014; Wen et al., 2008, Wen, Lee, Sun, Wang, Hu, et al., 2012).

The other experiment was in the northern part of Lake Taihu conducted at Meiliangwan (MLW, 31° 15' N, 120° 13' E) as part of the Taihu Eddy Flux Network (Lee et al., 2014) between August 2012 and September 2016. Lake Taihu is located in the Yangtze River Delta in Eastern China. The H<sub>2</sub>O, HDO, and H<sub>2</sub><sup>18</sup>O concentrations were measured at two heights (1.1 and 3.5 m) above the water surface with an IRIS water vapor isotope analyzer (Model 911-0004; Los Gatos Research, Mountain View, CA, USA) at 2 Hz. This analyzer was also modified to allow high sampling flow to improve its time response. Switching between the two intakes occurred every 30 s, and the measurement became stable in 5 s (about 10 datapoints; Xiao et al., 2017, their Figure 3). In this study, the last 15 datapoints, corresponding to the last 7.5 s, were used for calculations. The measurement site was located 250 m from the northern shore. To minimize land influence on the measurement, we restricted most of our analysis to the data collated in the wind direction sector of 140°–315°, corresponding to a fetch of 8–50 km. The only exception is the analysis of the  $\delta_E$  bias errors in relation to footprint where data collected under all wind directions were used (Section 3.3). The in situ calibration vapor was generated by a water vapor isotope standard source (Model 908-0003-9002; Los Gatos Research). The calibration was performed every 3 h. Each calibration cycle consisted of 5 concentrations and lasted for 30 min in total. The five calibration concentration values varied between different calibration cycles. Other details of this experiment can be found in Xiao et al. (2017).

Errors in the vapor isotope ratio measured by IRIS analyzers can arise from concentration dependence and from scale expansion or delta stretching, but with the former dominating the latter (Wen et al., 2008, Wen, Lee, Sun, Wang, Tang, et al., 2012). An ideal calibration strategy is to correct the measurement for both errors. However, cycling through vapor standards at multiple concentration and multiple delta values would take too much instrument time away from ambient measurement. At Zhangye and Lake Taihu, the calibration method deployed only one isotopic standard but at multiple concentration values, which aimed at removing the concentration dependence. A comparison between calibration using one delta standard versus that using multiple delta standards indicates that the one-delta calibration may introduce an error of about 0.1‰ (Wen, Lee, Sun, Wang, Tang, et al., 2012).

## 2.2. Regression Models

In this study, three regression models were used with the Keeling plot method to obtain the intercept of Equation 1: OLS, GMR, and YS. The OLS seeks to minimize the sum of the squared residuals between the

expected values of the dependent variable  $y$  and the data points. The result is unbiased only if errors in the independent variable  $x$  are negligible and errors in  $y$  are constant. The GMR seeks to minimize both the  $y$  and  $x$  residuals. The result is unbiased only when the normalized error in  $x$ , or error in  $x$  divided by the variance of  $x$ , is equal to the normalized error in  $y$  (Kermack & Haldane, 1950). In addition, the correlation between errors in  $x$  and in  $y$  is set to be zero for the OLS and the GMR methods. In other words, the OLS and GMR methods are two special cases of YS (Wehr & Saleska, 2017). The YS method is a general solution for the best fit line where observations are independent of each other, the errors in  $x$  and error in  $y$  are normally distributed, and these errors can be correlated (Wehr & Saleska, 2017; York, 1966, 1969; York et al., 2004). Other details on the regression models can be found in Chen et al. (2017) and Wehr and Saleska (2017).

### 2.3. Characterization of Error Structures

The error parameters in the YS model,  $\sigma(x_i)$ ,  $\sigma(y_i)$ , and  $r_i$ , for Lake Taihu were determined with the 2-Hz field calibration data. Here  $\sigma(x_i)$  and  $\sigma(y_i)$  are errors in horizontal ( $1/c$ ) and vertical ( $\delta_v$ ) coordinates at the  $i$ th datapoint, and  $r_i$  is correlation between errors in  $x_i$  and errors in  $y_i$ . A calibration cycle consisted of five concentrations, each lasting for 6 min. An example of the calibration stepping is given by Xiao et al. (2017; their Figure 4). The standard deviation of  $1/c$ , the standard deviation of  $\delta_v$  and the correlation coefficient between  $1/c$  and  $\delta_v$  were calculated for each concentration interval. We assumed that these variations originated purely from measurement errors. The data during transition from one concentration level to the next were excluded from the calculation.

For Zhangye, the instrument calibration was done in the field, but the field calibration data cannot be used to characterize the measurement errors because the concentration of the water vapor generated by the liquid vaporization module was not stable during the calibration phase. To obtain the error parameters, we carried out additional measurements after the field experiment had been completed, using the same IRIS water vapor isotope analyzer but a different calibration unit. The analyzer was configured to measure at 0.2 Hz the water vapor concentration and the isotopic composition of a water vapor stream generated by a standard delivery module (Model A0101; Picarro Inc.), which a different calibration unit than what was deployed in the field (Picarro Inc.). This new delivery module was fed with liquid water of known delta value ( $-9.17\%$ ). Each measurement cycle included three water vapor concentrations and lasted for 1 h. These concentration values varied between different measurement cycles. A total of 141 measurement cycles were performed, with the vapor concentration ranging from 7,900 to 27,690 ppm. This concentration range spanned the concentration variations during the field experiment. The same method used for Lake Taihu was used to calculate the error parameters.

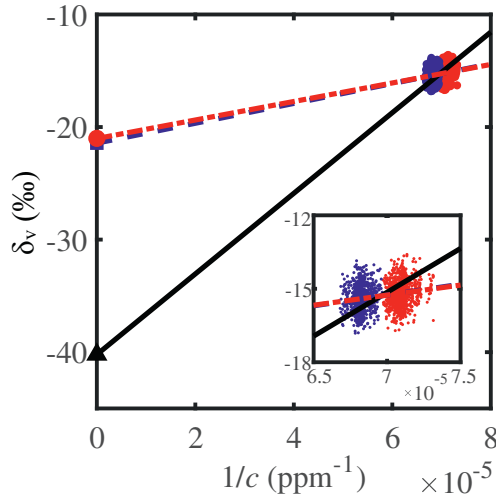
### 2.4. Data Processing

The high-frequency IRIS data were used to calculate  $\delta_E$  for each hourly observation interval. In the flux-gradient method, after each switching between the higher and lower intakes, the data were averaged to obtain the mean concentration differences between the two measurement heights, and  $\delta_E$  was determined from the gradient ratio according to Equation 2 after span correction as described by Lee et al. (2007). And then the hourly  $\delta_E$  and its corresponding standard deviation can be obtained.

In the Keeling plot method, the three regression models described above were used to determine  $\delta_E$ . Each observation, including data obtained for both measurement heights, consisted of about 200 and 1,800 data points for Zhangye and Lake Taihu, respectively. Figure 1 shows the time series of the water vapor mixing ratio and the calibrated vapor  $\delta$ , from a typical observation period at Lake Taihu, and Figure 2 shows the corresponding linear regression plot.

Three criteria were used to screen the data. The flux-gradient method becomes noisy at times of small vertical concentration gradients. To ensure a robust comparison, we restricted our analysis to observations whose hourly mean vertical vapor concentration difference between the two measurement heights was larger than 200 ppm in magnitude. About 2/3 of the 3,026 and 1,622 valid observations satisfy this criterion at Zhangye and Lake Taihu, respectively. The second criterion was the standard deviation of  $\delta_E$  calculated





**Figure 2.** An illustration of the three regression models applied to the data in Figure 1. YS: red-dot-dashed line,  $\delta_E = -21.02\text{‰}$ ; OLS: blue-dashed line,  $\delta_E = -21.43\text{‰}$ ; GMR: black-solid line,  $\delta_E = -40.21\text{‰}$ ; blue dots: observations at lower inlets; red dots: observations at higher inlet. In this hour, the  $\delta_E$  obtained from the flux-gradient method is  $-21.00\text{‰}$ , with a standard deviation of  $4.78\text{‰}$ . GMR, geometric mean regression; OLS, ordinary least squares; YS, York's solution.

by the flux-gradient method; we used a threshold value of  $20\text{‰}$ . The third criterion required that the  $P$  value obtained from the Keeling plot method be smaller than  $0.05$  to ensure that the relationship between  $1/c$  and  $\delta_v$  passes the significance test (Unger et al., 2010). A total of  $1,084$  and  $817$  hourly observations remained for Zhangye and Lake Taihu, respectively, after the three data screening criteria were applied.

## 2.5. CG Model of Lake $\delta_E$

The CG model was used to evaluate the bias errors of the Keeling plot and the flux-gradient methods for the lake site. The model computes  $\delta_E$  as,

$$\delta_E = \frac{\alpha_{\text{eq}}^{-1} \delta_L - h \delta_v - \epsilon_{\text{eq}} - (1-h) \epsilon_k}{1-h+0.001(1-h) \epsilon_k} \quad (3)$$

where  $\alpha_{\text{eq}} (>1)$  is the equilibrium fractionation factor which is a known function of water surface temperature (Majoube, 1971),  $\delta_L$  is the isotopic composition of the lake surface water,  $h$  is relative humidity in fraction,  $\epsilon_{\text{eq}} = 10^3 (1-1/\alpha_{\text{eq}})$  is the equilibrium factor in delta notation ( $\text{‰}$ ), and  $\epsilon_k$  is isotopic kinetic fractionation factor ( $\text{‰}$ ). The calculation was performed at hourly intervals using the measured variables as inputs. The kinetic factor was calculated with the wind-dependent parameterization of Merlivat and Jouzel (1979). In Merlivat and Jouzel (1979),  $\epsilon_k$  is an implicit

function of wind speed. Here we solved their function numerically and then fit the numerical solution with the following polynomial function,

$$\epsilon_k = 5.7818 \times 10^{-4} \times u^4 - 0.01825 \times u^3 + 0.2107 \times u^2 - 1.1428 \times u + 8.8617 \quad (4)$$

where  $\epsilon_k$  is in  $\text{‰}$  and in  $u$  is wind speed in  $\text{m s}^{-1}$ . This parameterization was independently validated by Xiao et al. (2017) against the measured local evaporation line and the isotopic mass balance of the lake. They found that the D and  $^{18}\text{O}$  isotopic compositions of evaporation calculated with this parameterization fall right on the line of the D and  $^{18}\text{O}$  compositions of the lake water.

## 2.6. Miller-Tans Method

The Miller-Tans method is also a popular method in the calculation of  $\delta_E$  (Miller & Tans, 2003). Different to the Equation 1, in the Miller-Tans method (Equation 5), the slope parameter is equivalent to the isotopic composition of surface evaporation, by establishing the relationship between  $\delta_v c$  and  $(c-c_0)$ ,

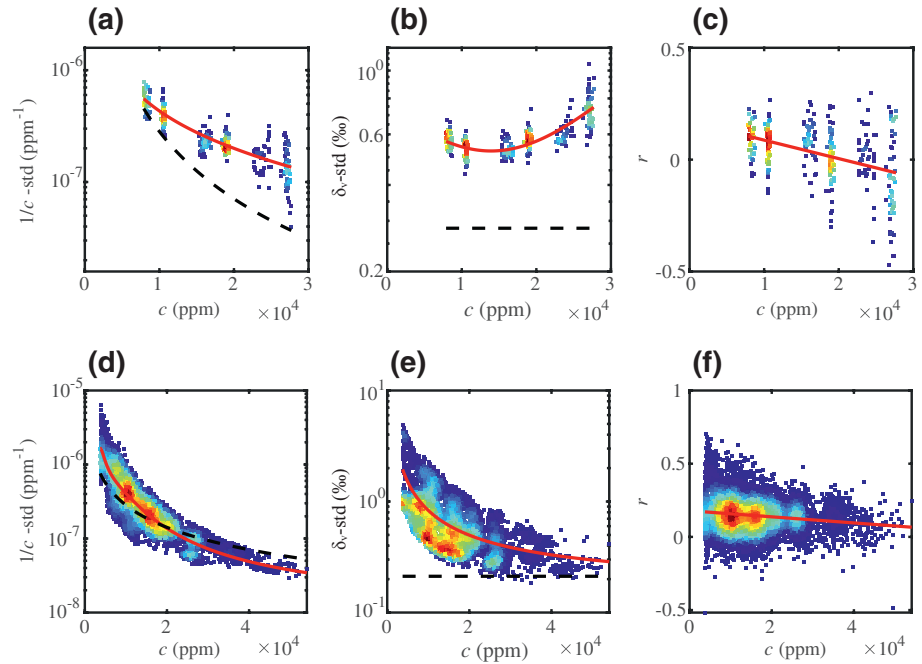
$$c \delta_v = d + a(c - c_0) \quad (5)$$

where  $a = \delta_E$  and  $d = \delta_0 c_0$  are slope and intercept coefficients, respectively.

## 3. Results and Discussion

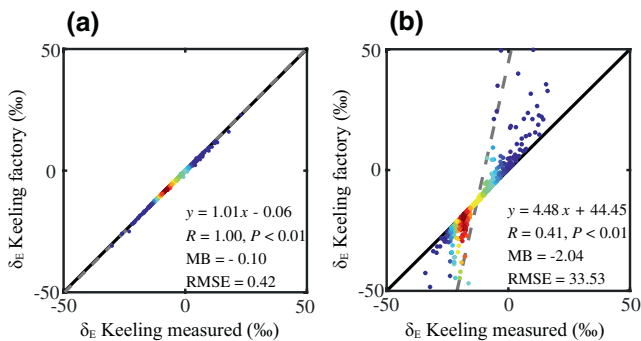
### 3.1. YS Model With Two Different Error Structures

The relationships between the water vapor concentration and the standard deviation of  $1/c$ , the standard deviation of  $\delta_v$  and correlation coefficient  $r$ , established with the data collected during the instrument calibration cycles, are shown in Figure 3. Unsurprisingly, the standard deviation of  $1/c$  was greater at lower concentrations, with the Picarro analyzer (used at Zhangye) and the LGR analyzer (used at Lake Taihu) giving similar performance. The standard deviation of  $1/c$  was  $4.02 \times 10^{-7}$  and  $4.07 \times 10^{-7} \text{ ppm}^{-1}$  at a water vapor concentration of  $10,000 \text{ ppm}$  and  $1.29 \times 10^{-7}$  and  $7.26 \times 10^{-8} \text{ ppm}^{-1}$  at a concentration of about



**Figure 3.** Relationships between water vapor concentration and errors in  $1/c$ , errors in  $\delta_v$  and correlation coefficient between errors in  $\delta_v$  and errors in  $1/c$  for Zhangye (a–c) and Lake Taihu (d–f). The solid red line and black dash line indicate the regression fit and the precision given by the manufacturers, respectively. The regression equations are given in Table S1. Color indicates data density. The data shown here are high-frequency (0.2 Hz at Zhangye and 2 Hz at Lake Taihu) calibration data.

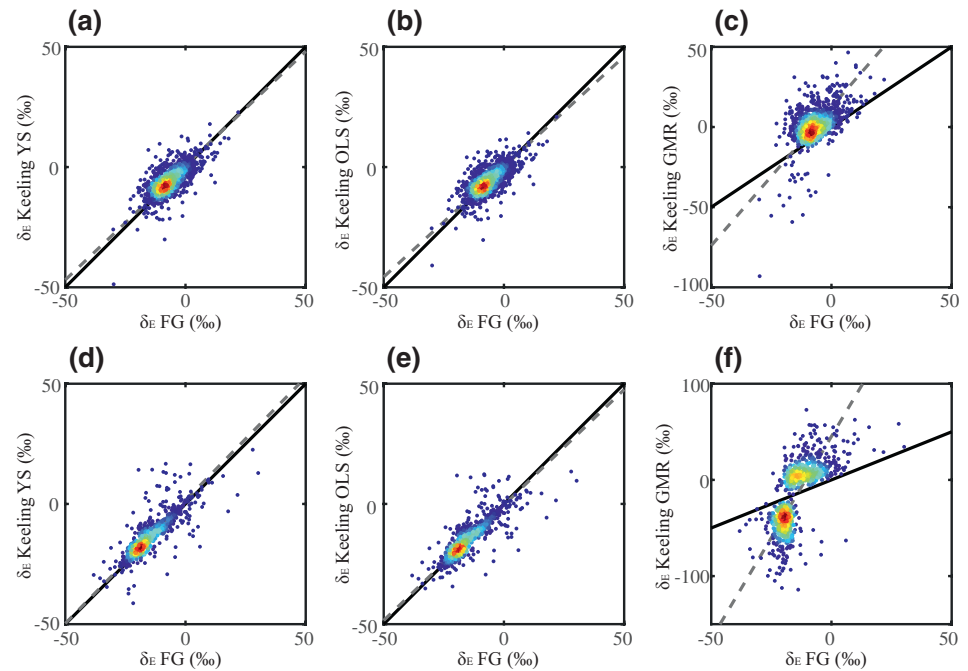
30,000 ppm, for Zhangye and Lake Taihu, respectively. The standard deviation of  $\delta_v$  showed opposite trends for the two sites. At Zhangye, the standard deviation of  $\delta_v$  was relatively constant at concentrations lower than about 20,000 ppm and increased slightly with increasing concentration beyond this threshold. At Lake Taihu, the standard deviation of  $\delta_v$  showed a general decreasing trend with increasing concentration. At Zhangye, the correlation between measurement errors in  $1/c$  and in  $\delta_v$  was slightly positive at low concentrations ( $\sim 10,000$  ppm) and varied around zero at high concentrations ( $\sim 25,000$  ppm). At Lake Taihu, the error correlation was mostly positive and did not seem to depend on the vapor concentration.



**Figure 4.** Comparison of the York's solution (YS) regression results with concentration-independent factory error structure and with concentration-dependent measured error structure for Zhangye (a) and Lake Taihu (b). Solid black lines are the 1:1 comparison and gray dash line is linear regression. The regression equation is shown in each panel along with the correlation coefficient  $R$ , mean bias (MB, ‰), and root mean square error (RMSE, ‰). In panel (b), 19 datapoints are out of the range of the vertical axis ( $\pm 50\%$ ). Color indicates data density.

For comparison, Figure 3 also shows error structures based on manufacturers' specifications. Specifically, the 0.1 Hz error (precision) is 20 ppm for water vapor and 0.20‰ for  $\delta^{18}\text{O}$  for Zhangye and the 1 Hz error is  $0.002c$  for water vapor and 0.15‰ for  $\delta^{18}\text{O}$  for Lake Taihu, and the correlation between the two variables is set to zero. The  $\delta^{18}\text{O}$  error is independent of concentration in both cases, and the  $1/c$  error can be expressed as  $20/c^2$  and  $0.002/c$  for Zhangye and Taihu, respectively. To match the observation frequency, the manufacturers' errors were multiplied by  $\sqrt{2}$  on the assumption that errors are inversely proportional to the square root of sample size. Generally, field errors were much larger than those specified by the manufacturers except for  $1/c$  at the high concentration range at Lake Taihu where the two were similar.

Results of regression fitting (Table S1) to the data shown in Figure 3 were used to determine parameters  $\sigma(x_i)$ ,  $\sigma(y_i)$ , and  $r_i$  in the YS regression model as functions of the measured concentration at observation  $i$ ,  $c_i$ . For example, the error in the vertical axis ( $\delta_v$ ) at time  $i$  is given as  $\sigma(y_i) = f_y(c_i)$ , where  $f_y$  is the regression fitting equation on  $c_i$ .



**Figure 5.** Comparison of the evaporation isotopic signature  $\delta_E$  obtained with the flux-gradient method and with the Keeling plot method for Zhangye (a–c) and Lake Taihu (d–f). Panels (a and d): YS regression model; panels (b and e): OLS regression model; panels (c and f): GMR regression model. Solid lines are the 1:1 comparison and dashed lines are linear regression. Refer to Tables 1 and 2 for regression statistics. Color indicates data density. GMR, geometric mean regression; OLS, ordinary least squares; YS, York's solution.

Figure 4 compares the YS calculation using the concentration-independent factory and concentration-dependent measured error structures. Nearly identical results were obtained for Zhangye using the two different error structures (panel a, linear correlation  $R = 1.00$ ,  $RMSE = 0.42\text{‰}$ ). This contrasts sharply with Lake Taihu, where the YS model with the manufacturer's error specification performed poorly against the YS estimate with field errors ( $R = 0.40$ ,  $RMSE = 32.60\text{‰}$ , Figure 4b), or against the OLS model ( $R = 0.42$ ,  $RMSE = 32.44\text{‰}$ ), the flux-gradient method ( $R = 0.33$ ,  $RMSE = 33.15\text{‰}$ ), and the CG model ( $R = 0.32$ ,  $RMSE = 34.29\text{‰}$ ; Figure S1). At the lake site, use of the concentration-dependent field error structures significantly improved the YS estimate in comparison with the OLS regression model ( $R = 0.99$ ,  $RMSE = 1.45\text{‰}$ ), the flux gradient method ( $R = 0.84$ ,  $RMSE = 4.60\text{‰}$ ) and the CG calculation ( $R = 0.65$ ,  $RMSE = 6.93\text{‰}$ ; Table 2). The disparity between the two YS estimates resulted from the fact the reported manufacturer's concentration-independent error for  $\delta_v$  is too small, but the actual error was higher and was also very sensitive to concentration (Figure 3; Salmon et al., 2019; Sturm & Knohl, 2010). The problem is particularly severe at the low concentration range, where according to the manufacturer, the normalized error in  $1/c$  is 0.018 and is only marginally better than the normalized error in  $\delta_v$  (0.029) whereas the actual normalized error in  $1/c$  (0.026) is much better than the actual normalized error in  $\delta_v$  (0.12; Table 3). When normalized errors in the dependent and independent variables are comparable, the YS model behaves like the GMR model (Wehr & Saleska, 2017).

In the following, we will restrict our discussion to the YS results using the concentration-dependent measured error structures.

## 3.2. Comparison Between the Flux-Gradient and the Keeling Plot Methods

### 3.2.1. Comparison Statistics

A comparison between the flux-gradient method and the Keeling plot method using the three regression models is summarized in Tables 1 and 2. Figure 5 shows the comparison in 1:1 plots, where each data



**Table 1**  
*Comparison Between the FG Method and the Keeling Plot Method With YS, OLS, and GMR Regression Models for Zhangye, Showing the Linear Regression, Correlation Coefficient (R), MB (‰), and RMSE (‰)*

		FG	Keeling with YS	Keeling with OLS
Keeling with YS	Equation	$y = 0.95x + 0.59$	–	–
	R	0.70		
	MB	0.93		
	RMSE	4.48		
Keeling with OLS	Equation	$y = 0.92x + 0.17$	$y = 0.97x - 0.41$	–
	R	0.70	1.00	
	MB	0.73	–0.20	
	RMSE	4.39	0.48	
Keeling with GMR	Equation	$y = 1.73x + 12.47$	$y = 1.82x - 11.40$	$y = 1.88x + 12.16$
	R	0.52	0.72	0.69
	MB	7.48	6.75	6.75
	RMSE	11.40	9.68	10.01

*Note.* The mean bias is calculated as the estimate using the method in the column header minus that using the method listed in the row header.

Abbreviations: FG, flux-gradient; GMR, geometric mean regression; MB, mean bias; OLS, ordinary least squares; RMSE, root mean square error; YS, York's solution.

point represents one hourly  $\delta_E$  calculated using data from two measurement heights above the surface. The Keeling plot results with the YS and OLS regression models and the flux-gradient results were comparable, with the relative mean biases (MBs) less than 1.01‰ among each other and high linear correlations ( $R > 0.70$ ) for both Zhangye and Lake Taihu.

In contrast, the GMR results were rather poor in comparison with the OLS, the YS or the flux-gradient method results. For example, the MB against the flux-gradient method was large, at 6.97 and  $-1.66\%$  for Zhangye and Lake Taihu, respectively. An implicit assumption of GMR is that the normalized errors in  $x$  and  $y$  are equal (Kermack & Haldane, 1950). This assumption was not satisfied here. Table 3 shows the mean errors in  $1/c$  and  $\delta_v$ , and these errors normalized by the ranges of  $1/c$  and  $\delta_v$  for three levels of water vapor concentration. Here the range of a variable is defined as the difference between the maximum and the minimum value of the high-frequency data in a given 60-min observational period. Errors existed both in the horizontal coordinate,  $1/c$ , and the vertical coordinate,  $\delta_v$ . The normalized error in  $\delta_v$  was, however, much larger than the normalized error in  $1/c$ , by a factor of 6–10 at Zhangye and of 4–6 at Lake Taihu.

Some researchers advocate for the Miller-Tans equation (Equation 5) instead of the Keeling plot equation (Equation 1) when using the GMR model. We used the OLS and the GMR models to estimate the slope parameter in Equation 5 and compared the results with the Keeling plot method using the OLS model (Figure S2). The Miller-Tans slope from OLS was nearly identical to the Keeling plot intercept from OLS. However, the Miller-Tans slope from GMR showed large deviations from the Keeling plot intercept from OLS. Comparison of the Miller-Tans method using GMR with estimates from Keeling plot method with YS, the flux-gradient method or the CG calculation yielded similarly large deviations. It appears that the GMR was not a good regression model for the high-frequency water vapor data deployed in this study, regardless of whether the Keeling plot equation or the Miller-Tans equation was used.

In CO<sub>2</sub> studies, increasing CO<sub>2</sub> concentration range will reduce the systematic bias associated with the  $\delta^{13}\text{C}$  signature of respired CO<sub>2</sub> inferred from the Keeling plot method with the OLS model. This is true for observations made with flasks (Pataki et al., 2003; Zobitz et al., 2006) and IRIS instruments (Chen et al., 2017; Zobitz et al., 2006) and for synthetic CO<sub>2</sub> datasets (Kaylor et al., 2010; Wehr & Saleska, 2017). In the case of water vapor, the performance of the OLS model, in reference to the flux-gradient method

**Table 2**  
*Comparison Among the FG Method, the CG Model Calculation and the Keeling Plot Method With YS, OLS, and GMR Regression Models for Lake Taihu, Showing the Linear Regression, Correlation Coefficient (R), MB (%), and RMSE (%)*

		CG	FG	Keeling with YS	Keeling with OLS
FG	Equation	$y = 0.99x + 0.18$	–	–	–
	R	0.72			
	MB	0.31			
	RMSE	5.96			
Keeling with YS	Equation	$y = 1.02x + 1.86$	$y = 1.00x + 1.07$	–	–
	R	0.65	0.84		
	MB	1.50	1.01		
	RMSE	6.93	4.60		
Keeling with OLS	Equation	$y = 0.96x - 0.10$	$y = 0.96x - 0.45$	$y = 0.96x - 1.47$	–
	R	0.66	0.83	0.99	
	MB	0.48	0.14	–0.87	
	RMSE	6.51	4.59	1.45	
Keeling with GMR	Equation	$y = 4.18x + 46.02$	$y = 4.21x + 45.24$	$y = 4.19x + 40.73$	$y = 4.39x + 47.20$
	R	0.53	0.68	0.79	0.80
	MB	–0.91	–1.66	–2.67	–1.80
	RMSE	30.18	28.64	27.64	27.72

The mean bias is calculated as the estimate using the method in the column header minus that using the method listed in the row header.

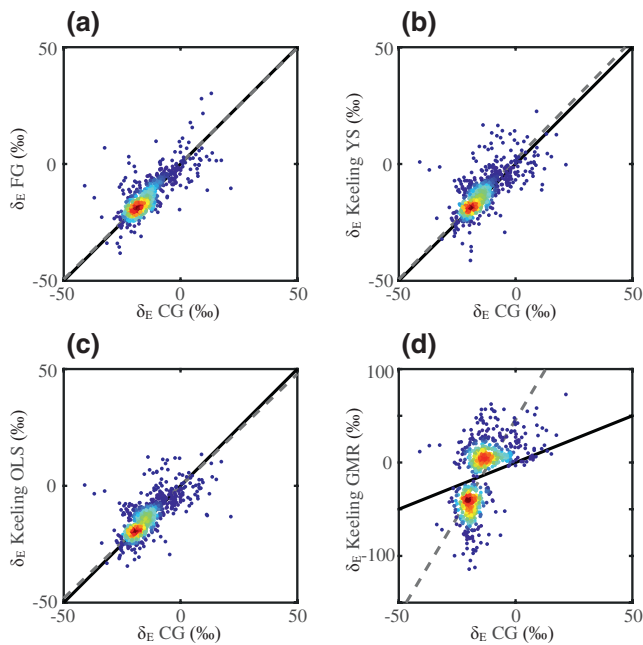
Abbreviations: CG, Craig-Gordon; FG, flux-gradient; GMR, geometric mean regression; MB, mean bias; OLS, ordinary least squares; RMSE, root mean square error; YS, York's solution.

or the YS estimate, did not show obvious dependence on concentration range (Figures S3 and S4). The lack of range dependence may be because the normalized error in  $\delta_v$  is about five times the normalized error in  $1/c$  at both the high and low concentration ranges (Table 3). In other words, the error structures assumed by the OLS method are a good approximation regardless of concentration range.

### 3.2.2. Results From Keeling Plot Method With Single-Height Data

Logistically, it is much easier to measure the water vapor isotopic composition at a single height than at multiple heights involving valve switching. Indeed, the great majority of the published IRIS water vapor isotope measurements have been conducted with single-height configurations (Fiorella et al., 2018; Wei et al., 2019; Yao et al., 2018; Zannoni, Steen-Larsen, Stenni, et al., 2019). Here, we applied the Keeling plot method to data collected at the lower and higher measurement height and compared the results with the flux-gradient method (Figures S5 and S6). All the three regression models performed less well when only data from the lower or higher height was used than when data from both heights were used to perform the regression (Figure 5). For example, at Lake Taihu, the mean difference of the OLS against the flux-gradient method changed to 0.44‰ (Figure S5e) from 0.14‰ (Figure 5e) and the correlation between the Keeling plot method and the flux-gradient method was reduced to 0.33 from 0.83.

The deteriorated performance in Figures S5 and S6 resulted mostly from a narrower concentration range which increase the uncertainty of the Keeling plot results (Chen et al., 2017; Pataki et al., 2003; Zobitz et al., 2006), with a reduced sample size being another (minor) factor. The sample size was halved when only one-height measurement was used. The mean concentration ranges were 2,063 and 1,137 ppm for Zhangye and Lake Taihu, respectively, for the data shown in Figure S5, whereas the mean ranges were larger, at 2,534 and 1,712 ppm for Zhangye and Lake Taihu, respectively, in Figure 5. Similarly, Good et al. (2012) reported that the uncertainty of the isotopic composition of evapotranspiration associat-



**Figure 6.** Comparison of the evaporation isotopic signature  $\delta_E$  against the Craig-Gordon (CG) model calculation for Lake Taihu: flux-gradient method (a), and Keeling plot method with the YS (b), the OLS (c), and the GMR (d) regression models. Solid lines are the 1:1 comparison and dashed lines are linear regression. Refer to Table 2 for regression statistics. Color indicates data density. GMR, geometric mean regression; OLS, ordinary least squares; YS, York's solution.

ed with high-frequency time series measured at a single height is 25% larger than with a combined use of time series measured at multiple heights.

While the systematic biases of the GMR regression in Figures S5 and S6 (panels c and f) are mathematical in nature (as in Figure 5 panels c and f), the biases of the OLS and the YS regression here may be related to footprint influences, especially at Zhangye where the fetch was short (about 200 m). Griffis et al. (2007) found that the Keeling plot method with one-height data yields lower estimates of the  $\delta^{13}\text{C}$  composition of ecosystem respiration of a  $\text{C}_4$  crop than the flux-gradient method. They attributed this difference to a footprint mismatch: the single-height concentration has a much larger source area and is therefore more influenced by the surrounding  $\text{C}_3$  crops, than the flux-gradient data. Interestingly, Good et al. (2012) also found higher  $\delta_E$  values with the Keeling plot method using single-height data than using data from multiple heights (mean difference about 16‰ for  $^{18}\text{O}$ , black triangles in their Figure 8).

### 3.3. Comparison With the CG Model

It is believed that the CG model accurately predicts the  $\delta_E$  of evaporating water bodies, and we tested the  $\delta_E$  determined with the flux-gradient method and the Keeling plot method using each of the three regression models at Lake Taihu against this benchmark (Figure 6). The comparison statistics are summarized in Table 2. Once again, the GMR result showed a very large RMSE (30.18‰). The flux-gradient method and the Keeling plot method with the OLS and the YS regression were comparable in terms of linear correlation and RMSE. Of these three estimates, the flux-gradient and the OLS regression showed a small MB of 0.31‰ and 0.48‰, respectively, and the MB of the YS regression was larger, at 1.50‰.

The relative larger bias between YS and CG begs the question of whether the CG model of isotopic evaporation can be used to benchmark the performance of the flux-gradient method or the Keeling plot method. For a given hourly period, measurement errors in the CG input variables can propagate through the model to cause errors in the calculated  $\delta_E$ . This type of error should be random. One potential source of systematic CG error lies in the parameterization of the kinetic fractionation of evaporation. In the present study, we used the wind dependent parameterization of Merlivat and Jouzel (1979), which has been validated independently against the observed local evaporation line and the observed lake evaporation (Xiao et al., 2017). The mean kinetic factor for the observations shown in Figure 6 was 7.3‰. Forcing the CG model to remove the MB with the YS result would require that the mean kinetic factor be lowered to 5.8‰. Such a small kinetic factor seems unphysical because it is lower than other values used in the literature for inland water bodies (Gonfiantini et al., 2018; Jasechko et al., 2014) and is even lower than that for the open ocean (Merlivat & Jouzel, 1979). Instead, we suggest that the relative bias between YS and CG stemmed from the error structure used for YS. The instrument errors were highly sensitive to concentration. The error structure shown in Figure 3 was based on measurement taken during instrument calibration cycles. While it was much better than the error structure based on the instrument specification, it may still deviate from the true measurement errors.

The results presented above were based on open-fetch conditions. Here, we investigated the relationship of the bias errors to flux footprint. The flux footprint model of Kljun et al. (2015) was used to determine the fractional contribution from the land surface for each hourly observation (Figure S7). The main input variables (measurement height, friction velocity, the Obukhov length, standard deviation of the cross-wind component, wind direction, and wind speed) were provided by our measurement system. The roughness length was set to  $8.04 \times 10^{-4}$  m according to the locally tuned momentum transfer coefficient (Xiao et al., 2013). The boundary layer height was provided by the Global Data Assimilation System of the US National Oce-

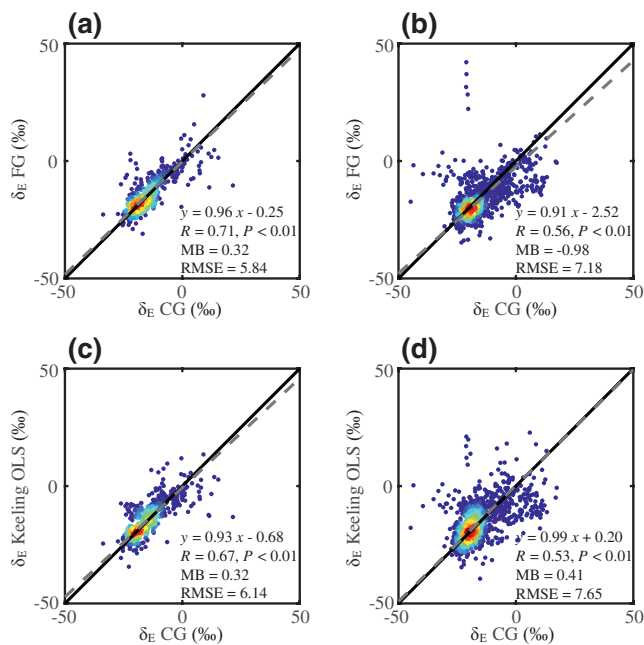
**Table 3**  
Measurement Errors in Three Quantiles of Water Vapor Concentration

Quantile	Mean $c$ (ppm)	Error in $1/c$ (ppm <sup>-1</sup> )	Error in $\delta_v$ (‰)	$c$ range (ppm)	$\delta_v$ range (‰)	$1/c$ range (ppm <sup>-1</sup> )	Error in $1/c/(1/c$ range)	Error in $\delta_v/$ ( $\delta_v$ range)
Zhangye								
0–25	7,688	$5.86 \times 10^{-7}$ ( $4.81 \times 10^{-7}$ )	0.57 (0.28)	1,856	3.48	$3.64 \times 10^{-5}$	0.016 (0.013)	0.16 (0.081)
25–75	12,984	$3.32 \times 10^{-7}$ ( $1.68 \times 10^{-7}$ )	0.53 (0.28)	2,659	3.50	$1.65 \times 10^{-5}$	0.020 (0.010)	0.15 (0.081)
75–100	18,930	$2.16 \times 10^{-7}$ ( $7.89 \times 10^{-8}$ )	0.55 (0.28)	3,566	3.95	$1.02 \times 10^{-5}$	0.021 (0.008)	0.14 (0.072)
Lake Taihu								
0–25	9,171	$4.58 \times 10^{-7}$ ( $3.08 \times 10^{-7}$ )	0.88 (0.21)	1,262	7.43	$1.73 \times 10^{-5}$	0.026 (0.018)	0.12 (0.029)
25–75	21,213	$1.25 \times 10^{-7}$ ( $1.33 \times 10^{-7}$ )	0.49 (0.21)	1,697	3.29	$4.27 \times 10^{-6}$	0.029 (0.031)	0.15 (0.065)
75–100	30,879	$6.96 \times 10^{-8}$ ( $9.16 \times 10^{-8}$ )	0.40 (0.21)	2,232	3.11	$2.35 \times 10^{-6}$	0.030 (0.039)	0.13 (0.068)

Notes. Errors are calculated as one standard deviation of high frequency data (0.2 Hz at Zhangye and 2 Hz at Lake Taihu). Here  $c$  and  $\delta_v$  denote water vapor concentration and vapor  $\delta_v$ , respectively. Numbers in parentheses are using error estimates from instrument manufacturers.

anic and Atmospheric Administration (<https://ready.arl.noaa.gov/gdas1.php>). The fractional contribution by the land surface to the total flux varied from zero to about 12% (Figure S8). Figure 7 compares the flux-gradient and the Keeling method (with OLS) against the CG model for observations made under two flux footprint conditions: those with “clean” footprint (panels a and c; fractional land contribution less than 1%) and those with “contaminated” footprint (panels b and d; fractional land contribution greater than 3%). At land fractional contributions less than 1%, both the flux-gradient method and the Keeling plot with OLS

were in good agreement with the CG model (MB 0.32‰), consistent with the assessment based on the data collected under open-fetch conditions (Figure 6). During periods when the land source contribution exceeded 3%, the flux-gradient method showed a larger magnitude of MB of 0.98‰ than the Keeling method (0.41‰). It appears that, by also incorporating time variations caused by atmospheric turbulence within each hourly observation, the Keeling method was less sensitive to footprint influence than the flux-gradient method which relied on the hourly mean concentration values only for the determination of  $\delta_E$ .



**Figure 7.** Comparison of the evaporation isotopic signature  $\delta_E$  against the Craig-Gordon (CG) model calculation for Lake Taihu: the flux-gradient method (a and b) and the Keeling plot method with OLS (c and d) under different footprint conditions. Panels (a and c) are for observations with “clean” footprint (fractional contribution from land less than 0.01), and panels (b and d) are for observations with “contaminated” footprint (fractional contribution from land greater than 0.03). The solid line is 1:1 and the dashed line is the linear regression with regression statistics noted. MB, mean bias; OLS, ordinary least squares; RMSE, root mean square error.

#### 4. Summary and Conclusions

Implementation of Keeling plot method requires that measurement errors in the concentration and in the isotopic composition be known accurately. The OLS and GMR regressions assume extreme error structures: In OLS, error is zero for  $x$  and is equal across the data range for  $y$ ; in GMR, the normalized error in  $x$  and  $y$  are equal. In principle, YS is the best choice of the three regression models if measurement errors in the concentration and in the isotopic composition and their correlations are known (Wehr & Saleska, 2017). Wehr and Saleska (2017) found that the YS regression with concentration-independent errors yields the least biased result for  $\text{CO}_2$  isotopes. In the present study, the YS regression with concentration-independent (factory-specified) isotope errors worked well for one instrument (Picarro at Zhangye) but not for the other (LGR at Lake Taihu). Use of the concentration-dependent error structures from field observations significantly improved the YS estimate in comparison with the OLS regression, the flux-gradient method and the CG model calculation. This difference resulted mainly from the fact that the manufacturer default error value for  $\delta_v$  is invariant with the vapor concentration whereas the actual error was highly dependent on the concentration.

The Keeling plot results with the YS (using field error structures) and OLS regression and the flux-gradient results were comparable, with the mean difference in the hourly  $\delta_E$  of less than 1.01‰ and with high linear correlation ( $R = 0.70$ – $1.00$ ) for both Zhangye and Lake Taihu. These results were obtained with high frequency measurements made at two heights above the surface. The agreement with the flux-gradient method deteriorated ( $R = 0.29$ – $0.49$ ) if one-height data was used to perform the YS and the OLS regression. In general, the GMR results were poor in comparison with the OLS and the YS regression or with the flux-gradient method. Unlike in studies of carbon isotopes, here the GMR bias in reference to the OLS regression could be either positive or negative. Use of the Miller-Tans instead of the Keeling plot equation did not bring improvement to the GMR performance.

At Lake Taihu, the flux-gradient method and the Keeling plot method with the OLS and the YS regression (using field error structures) were comparable to the CG model calculation in terms of linear correlation  $R$  ( $0.66$ – $0.72$ ) and RMSE ( $5.96$ – $6.93$ ‰). The MB error in  $\delta_E$  was small for the flux-gradient method ( $0.31$ ‰) and the Keeling plot method with the OLS regression ( $0.48$ ‰), indicating that these methods were both robust, at least for this site with large fetch conditions. The MB of the YS regression was larger ( $1.50$ ‰), suggesting room for further improvement of the error structures used for YS. Compared with the periods when the land source contribution was smaller than 1%, both the flux-gradient method and the Keeling plot method showed a larger magnitude of MB of  $0.98$ ‰ and  $0.41$ ‰, respectively, when the land source contribution exceeded 3%. The Keeling method with OLS regression may be less sensitive to fetch than the flux-gradient method.

## Data Availability Statement

The water vapor isotope data used in this study are available on the website <https://vapor-isotope.yale.edu>.

## Acknowledgments

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